

A variance reduction method for computing VaR

- 1. Computing Value at Risk by Monte Carlo simulations**
- 2. Importance Sampling for variance reduction**
- 3. Interacting Particle Systems for Importance Sampling (IPS-IS)**
- 4. Simulation results**

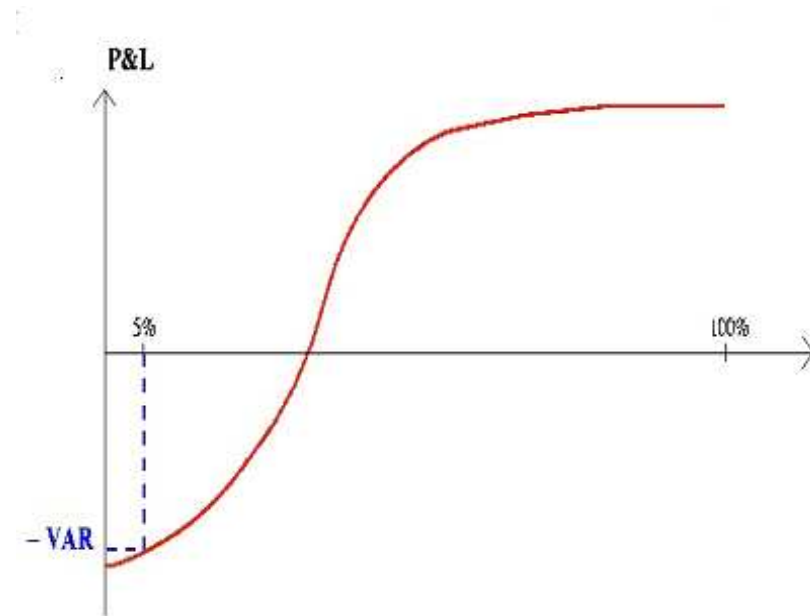
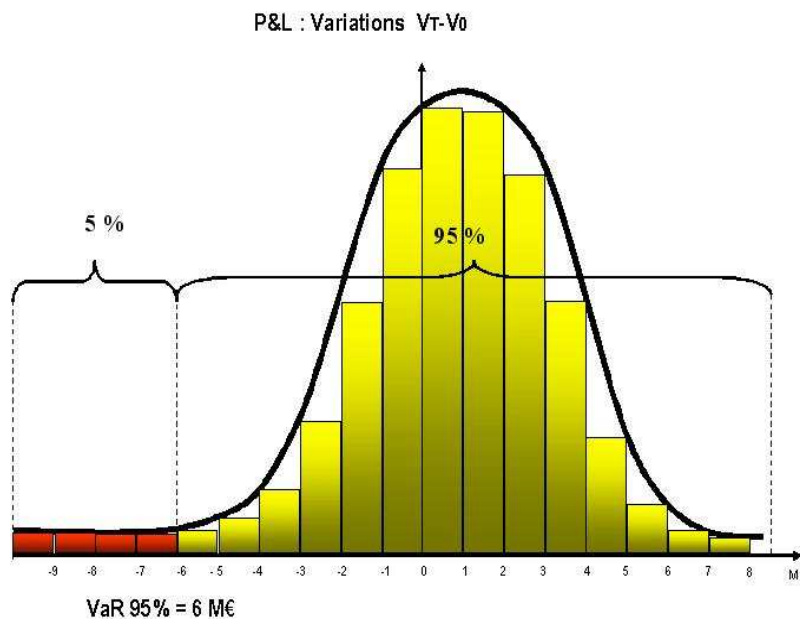
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Value at Risk and quantile

► **P&L of a portfolio on $[0, T]$** $\Delta V(X) = V_T - V_0 + \int_0^T CF$

with $X \in \mathbb{R}^d$ the risk factors impacting the portfolio value on $[0, T]$

► **Value at Risk** $VaR_\alpha = | \inf\{s \in \mathbb{R} \mid \mathbb{P}(\Delta V \leq s) \geq 1 - \alpha\} |$



$$VaR_\alpha = |F^{-1}(\alpha)|, \quad \text{where } F(s) = \mathbb{P}(\Delta V \leq s), \quad \text{for all } s \in \mathbb{R}$$

Monte Carlo method for VaR estimation

- ▶ The distribution function F can be viewed as an expectation

$$F(s) = \mathbb{E}[\mathbf{I}_{\Delta V(X) \leq s}], \quad \text{for all } s \in \mathbb{R}$$

- ▶ Traditional Monte Carlo Method for computing VaR

1. Monte Carlo simulations give an approximation of $F(s)$:

$$\hat{F}_N(s) = \frac{1}{N} \sum_{i=1}^N \mathbf{I}(\Delta V(X_i) \leq s), \quad \text{for all } s \in \mathbb{R}$$

⇒ Too many evaluations of ΔV for a given accuracy

2. Inversion of \hat{F}^N and interpolation for approximating VaR

Importance Sampling for variance reduction

- **Change of measure** $p \longrightarrow q$ where q dominates Hp

$$m = \mathbb{E}_p[H(X)] = \mathbb{E}_q\left[H(Y) \frac{p}{q}(Y)\right], \quad \text{where } X \sim p \text{ and } Y \sim q$$

- **Optimal change of measure** $p \longrightarrow q^*$ achieves **zero variance** if $H \geq 0$

$$q^* = \frac{Hp}{\int H(x)p(x) dx} = \frac{Hp}{\mathbb{E}_p[H(X)]} = H \cdot p$$

- **Monte Carlo approximation**

$$\mathbb{E}_p[H(X)] \approx \hat{m}_M^q = \frac{1}{M} \sum_{i=1}^M H(Y_i) \frac{p}{q}(Y_i), \quad \text{where } (Y_1, \dots, Y_M) \text{ i.i.d. } \sim q$$

⇒ **How to simulate and evaluate approximately q^* ?**

Variance of the Importance Sampling estimate

- Let q be a (possibly random) importance probability density dominating q^*

$$\text{Var}(\hat{m}_M^q) = \mathbb{E} \left[\text{Var}[\hat{m}_M^q | \mathcal{F}_q] \right] + \underbrace{\text{Var} \left[\mathbb{E}[\hat{m}_M^q | \mathcal{F}_q] \right]}_{=0}$$

\mathcal{F}_q denotes the sigma-algebra generated by the random variables involved in q

- The variance of the IS estimate depends on the "distance" between q and q^*

$$\text{Var}(\hat{m}_M^q) = \frac{m^2}{M} \mathbb{E} \left[\int \left[(q^* - q) \frac{q^*}{q} \right] (x) dx \right]$$

- **Idea** : use **Interacting Particle Systems for Importance Sampling (IPS-IS)** to approximate q^* by q^N based on an N -particle system to achieve

$$\text{Var}(\hat{m}_M^{q^N}) \leq \frac{C}{MN^\alpha} \quad \text{with} \quad 0 < \alpha < 1/2$$

Some alternative approaches

▶ **Large deviation approximation** for rare events simulation

▶ **Approximation of H** to obtain a simple form for q^*

ex : [Glasserman&al00] for computing VaR, Δ - Γ approximation of the portfolio

▶ **Cross-entropy** [Homem-de-Mello&Rubinstein02]

q^θ is chosen in a parametric family such as to minimize the entropy $K(q^\theta, q^*)$

▶ **Interacting Particle Systems** without Importance Sampling

[DelMoral&Garnier05], [Cerou&al06]

Interacting Particle Systems for Importance Sampling (IPS-IS) can be viewed as a non parametric version of cross entropy approach

Progressive correction [Musso&al01]

- We introduce a sequence of non negative functions $(G_k)_{0 \leq k \leq n}$ such that

$$\text{for all } x \in \mathbb{R}^d, \quad \left\{ \begin{array}{l} G_0(x) = 1 \\ \text{The product } G_0(x) \cdots G_n(x) = H(x) \\ \text{If } G_k(x) = 0 \text{ then } G_{k+1}(x) = 0 \end{array} \right.$$

- In our case $H(x) = \mathbf{I}(\Delta V(x) \leq s)$ then we choose

$$G_k(x) = \mathbf{I}_{\Delta V(x) \leq s_k}, \quad \text{with } s = s_n \leq \cdots \leq s_0 = +\infty$$

- Dynamical system on the space of probability measures $(\nu_k)_{0 \leq k \leq n}$

$$\left\{ \begin{array}{l} \nu_0 = p dx \\ \nu_k = \frac{G_k \nu_{k-1}}{\int_{\mathbb{R}^d} G_k(x) \nu_{k-1}(x) dx} = G_k \cdot \nu_{k-1}, \quad \text{for all } 1 \leq k \leq n \end{array} \right.$$

$$\Rightarrow \nu_n = q^* dx$$

Space exploration

- We introduce a sequence of Markov kernels $(Q_k)_{0 \leq k \leq n}$ such that

$$\nu_k \approx \nu_k Q_k \quad \text{i.e.} \quad \nu_k(dx) \approx \int_{\mathbb{R}^d} \nu_k(du) Q_k(u, dx), \quad \text{for all } x \in \mathbb{R}^d$$

- In our case where $G_k(x) = \mathbf{I}_{\Delta V(x) \leq s_k}$, if p is Gaussian then Q_k is easily obtained from a **Gaussian kernel** Q reversible for p ,

$$Q_k(x, dx') = Q(x, dx') \mathbf{I}_{\Delta V(x) \leq s_k} + [1 - Q(x, \Delta V^{-}((-\infty, s_k]))] \delta_x(dx')$$

- Dynamical system on the space of probability measures $(\nu_k)_{0 \leq k \leq n}$

$$\begin{cases} \nu_0 = p dx \\ \nu_k = G_k \cdot (\nu_{k-1} Q_{k-1}), \quad \text{for all } 1 \leq k \leq n \end{cases}$$

$$\Rightarrow \nu_n = q^* dx$$

Approximation of the dynamical system

► The idea is to replace at each iteration k , $\nu_{k-1} Q_{k-1}$ by its N -empirical measure $S^N(\nu_{k-1} Q_{k-1})$ such that

$$S^N(\nu_{k-1} Q_{k-1}) = \frac{1}{N} \sum_{i=1}^N \delta_{X_k^i} \quad \text{where } (X_k^1, \dots, X_k^N) \text{ are i.i.d. } \sim \nu_{k-1} Q_{k-1}$$

► Dynamical system on the space of discrete probability measures $(\nu_k^N)_{0 \leq k \leq n}$

$$\begin{cases} \nu_0^N = S^N(\nu_0) \\ \nu_k^N = G_k \cdot S^N(\nu_{k-1}^N Q_{k-1}), \quad \text{for all } 1 \leq k \leq n \end{cases}$$

⇒ One can show that $\nu_n^N \approx q^* dx$ [DelMoral]

Algorithm

- **Initialization** : Generate independently

$$(X_0^1, \dots, X_0^N) \text{ i.i.d. } \sim p \text{ then set } \nu_0^N = \frac{1}{N} \sum_{i=1}^N \delta_{X_0^i}$$

- **Selection** : Generate independently

$$(\tilde{X}_k^1, \dots, \tilde{X}_k^N) \text{ i.i.d. } \sim \nu_k^N = \sum_{i=1}^N \omega_k^i \delta_{X_k^i}$$

- **Mutation** : Generate independently for each $i \in \{1, \dots, N\}$,

$$X_{k+1}^i \sim Q_k(\tilde{X}_k^i, \cdot)$$

- **Weighting** : For each particle $i \in \{1, \dots, N\}$, compute

$$\omega_{k+1}^i = \frac{G_{k+1}(X_{k+1}^i)}{\sum_{j=1}^N G_{k+1}(X_{k+1}^j)} \text{ then set } \nu_{k+1}^N = \sum_{i=1}^N \omega_{k+1}^i \delta_{X_{k+1}^i}$$

Adaptive choice of the sequence $(G_k)_{0 \leq k \leq n}$

[Musso&al01], [Hommem-de-Mello&Rubinstein02], [Cérou&al06]

► The performance of Interacting particle systems is known to **deteriorate**

when the quantities $\frac{\max G_k}{S^N(\nu_{k-1}^N Q_{k-1})(G_k)}$ are big

The idea is then to chose G_k such that $\frac{1}{N} \sum_{i=1}^N G_k(X_k^i)$ is not to small

► In our case where $G_k(x) = \mathbf{I}_{\Delta V(x) \leq s_k}$, the threshold s_k is chosen as a r.v. depending on the current particle system and on a parameter $\rho \in (0, 1)$:

$$s_k = \inf \left\{ s \text{ such that } \sum_{i=1}^N \mathbf{I}_{\Delta V(X^i) \leq s} \geq \rho N \right\}$$

► This choice of s_k is not proved to guarantee that the algorithms ends in a finite number of iterations but this point does not seem to be a problem in our simulations

Density estimation

► At the end of the algorithm, we get $\nu_n^N \approx q^* dx$

But Importance Sampling requires a smooth approximation with density q^N

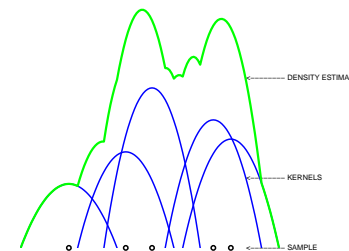
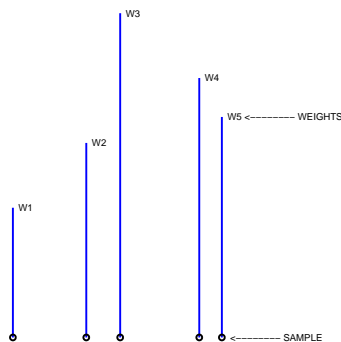
► **Kernel of order 2** K

$$K \geq 0 \quad \int K = 1 \quad \int x_i K = 0 \quad \int |x_i x_j| K < \infty$$

► **Rescaled kernel** K_h
$$K_h(x) = \frac{1}{h^d} K\left(\frac{x}{h}\right)$$

►
$$\nu^N = \sum \omega^i \delta_{X^i} \xrightarrow[\text{Density estimation}]{K_h * \cdot} q^{N,h} = \sum \omega^i K_h(\cdot - X^i)$$

► **Optimal choice of h** $\Rightarrow \mathbb{E} \|q^N - q^*\|_1 \leq \frac{C}{N^{\frac{4}{2(d+4)}}$



Some simulation results : Variance ratio

- ▶ Several test cases depending on the form of function $x \mapsto \Delta V(x)$ have been studied : results are all comparable
- ▶ X is a d dimensional Gaussian variable and $m = \mathbb{E}_p[\mathbf{I}_{\Delta V(X) \leq s}]$
- ▶ Particles $N = 500$ Iterations $n \approx 10$ to 60 Simulations $M = 10\,000$

	$d = 1$	$d = 2$	$d = 3$	$d = 4$	$d = 5$
$m = 10^{-2}$	150 10^{-1}	50	50	30	25
$m = 10^{-3}$	1000 2	300	300	200	140
$m = 10^{-6}$	$2 \cdot 10^5$ 200	10^5 400	10^5 300	$5 \cdot 10^4$ 460	$2 \cdot 10^4$ 480

	$d = 6$	$d = 7$	$d = 8$	$d = 9$...	$d = 30$
$m = 10^{-2}$	22	14	11	8	...	$5 \cdot 10^{-3}$
$m = 10^{-3}$	100	70	55	40	...	10^{-3}
$m = 10^{-6}$	10^4 250	$2 \cdot 10^3$ 480	$2 \cdot 10^3$ 300	$4 \cdot 10^3$ 300	...	1 360

References

- ▶ **[Glasserman&al00]** Glasserman, P. and Heidelberger, P. and Shahabuddin, P. *Reduction Variance techniques for estimating Value at Risk*, *Management Science*, Vol. 46, No 10, 2000.
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